

First-passage time analysis for Markovian deteriorating model

Gejza Dohnal

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First-affect time

Renewal period

Conclusions

Successive events and Energy Networks Markovian deteriorating model

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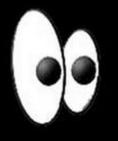
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August 14th, 2003.

That day in the northeastern US the temperature reached 90°C, which is more than the average of 80°F but less than the maximum 100°F. In the network, power flows dominated from the south and west to the north (Michigan) and East (New York). Export from the south and west of the US directed mainly to northern Ohio to Michigan and Ontario, Canada. The fault occured in northern Ohio due to a short circuit on contact wires overhead lines of 345 kV with a tree. Overloading caused the cascading spread of successive faults, gradual shutdown of transmission lines due to protection actions, formation of island traffic and then blackout in large parts of the northeast United States and southeast Canada. The total loss in sixteen hours accounted for 61,800 MW. The event hit 265 plants with 508 blocks. Without electricity remained 50 million inhabitants.





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Failure	×	Disaster	



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Failure	×	Disaster
random occurrence in time		random occurrence in time
independent recurrence		growing spread
renewal process		domino effect
quick repair or renewal	1	epairs slower than spreading



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Failure×Disasterrandom occurrence in time
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repairs slower than spreading

equipment failure accidents, health damage information system failure motor vehicle accident natural accident

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explosion, fire disease, epidemic system breakdown, piracy attack traffic collapse natural disaster

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 $\textbf{Model} \rightarrow \textbf{Prediction} \rightarrow \textbf{Prevention}$



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 $\textbf{Model} \rightarrow \textbf{Prediction} \rightarrow \textbf{Prevention}$

Preventive Maintenance Policy

Disaster Recovery Plan



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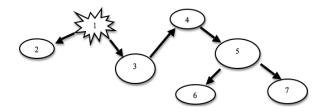
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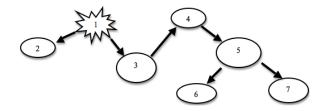
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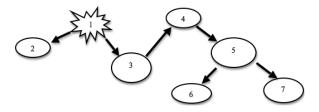
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- (i) at the beginning (at the time 0), the system is in the state $\omega^0 = (0, 0, \dots, 0)$,
- (ii) the process starts by deterioration of an object *i* with probability π_i , i = 1, 2, ..., n,
- (iii) when at the time t an object i was affected, there was a random time period τ after which the event moved onto some of the unaffected objects,
- (iv) states of the system in time *t* create a stochastic process $\{X(t), t \ge 0\}$ in continuous time. Values of this process lie within the set $\Omega = \{0, 1\}^n$



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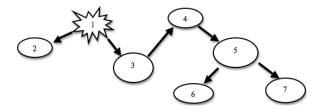
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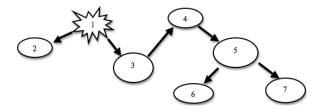
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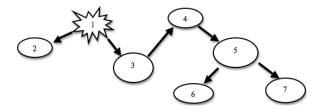
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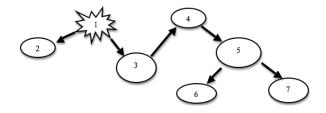
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Example (n = 7 $\pi = (1, 0, 0, 0, 0, 0, 0)$)

From all 2⁷ possible states only the following 14 are admissible:

 $\begin{array}{l} (1,0,0,0,0,0), (1,1,0,0,0,0), (1,0,1,0,0,0,0), (1,1,1,0,0,0,0), \\ (1,0,1,1,0,0,0), (1,1,1,1,0,0,0), (1,0,1,1,1,0,0), (1,1,1,1,1,0,0), \\ (1,0,1,1,1,1,0), (1,0,1,1,1,0,1), (1,1,1,1,1,1,0), (1,1,1,1,1,0,1), \\ (1,0,1,1,1,1,1), (1,1,1,1,1,1). \end{array}$



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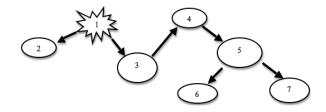
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Model 1 – One-way Model



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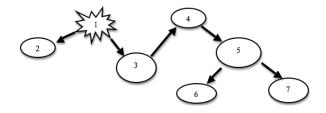
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Model 1 – One-way Model

Let us consider the following supplementary assumptions:

(v) an event can only affect one object in one moment,

- (vi) an event can only occur once on a particular object,
- (vii) the process moves to the next object with a probability which depends only on the recent state, not on the path leading to the recent state (the time sequence of events).



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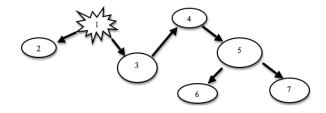
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Model 1 – One-way Model

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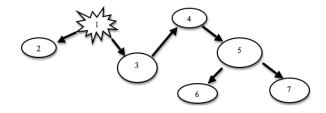
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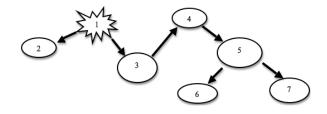
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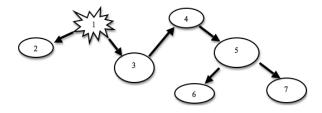
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Let us denote:

■ $|\omega| = \omega_1 + \omega_2 + \cdots + \omega_n$ the number of objects, on which a disastrous event holds when system is in the state ω (*size of deterioration* of the system)



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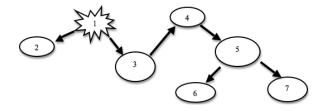
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- $\Omega_j = \{\omega \in \Omega : |\omega| = j\}$ contains all states, in which the pursued event hold exactly on *j* objects. The number of such states equals to $\binom{n}{j}$.



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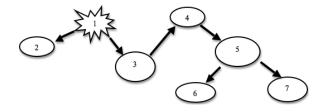
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$$\Omega = \bigcup_{j=1}^{n} \Omega_j, \ \ \Omega_i \cap \Omega_j = \emptyset \ \ \text{for all} \ \ i \neq j$$



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Proposition

Let us order states of the system in ascending order by the size of deterioration. Then, under assumptions (i)-(vii), the transition intensity matrix Q of the process $\{X(t), t \ge 0\}$ is upper triangular of the size $2^n \times 2^n$. The matrix Q can be written in the block-form as

	$(D_{0,0})$	P _{0,1}	O _{0,2}	O _{0,3}	•••	0)
	O _{1,0}	$D_{1,1}$	P _{1,2}	O _{1,3}	•••	O _{1,n}
Q =	O _{2,0}	O _{2,1}	D _{2,2}	P _{2,3}	•••	O _{2,n}
	:	÷	:	:	·	:
) 0	O _{n,1}	O _{n,2}	O _{n,3}	•••	0 0 _{1,n} 0 _{2,n} : 0

where $P_{i,j}$ is a rectangular matrix of size $\binom{n}{i} \times \binom{n}{j}$, the symbol O_{ij} denotes a null matrix and D_{ii} are square diagonal matrices for i = 1, ..., n. Moreover,

$$-D_{i,i} = \mathbf{e}.P_{i,i+1}'.I_{i,i}$$

where $I_{i,i}$ is the unit matrix of size *i*, *e* is the row vector of $\binom{n}{i}$ ones.



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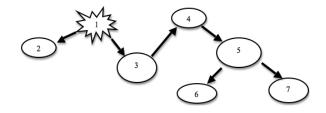
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Model 2 – Repairable Model

Let us omit the assumption (vi):

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 - ⇒ States of the system in time *t* create a Markovian process ${X(t), t \ge 0}$ in continuous time.



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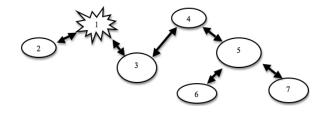
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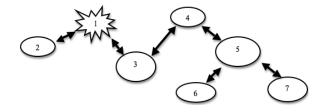
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The process X(t) can be considered as a random walk between sets $\Omega_0, \ldots, \Omega_n$. When the process in a state $\omega \in \Omega_j$, 0 < j < n, the only transitions to some states in Ω_{j-1} or Ω_{j+1} are allowed, whereas $\Omega_0 = \{\omega^0\}$ and $\Omega_N = \{\omega^N\}$ are reflection states. In the model, all states are transient. The process can be understood as the process of event spreading in the system with repair.



Model 2 – Repairable Model

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Proposition

In the model 2, the transition intensity matrix Q of the process $\{X(t), t \ge 0\}$ has the following block-form

	$(D_{0,0})$	P _{0,1}	<i>O</i> _{0,2}	O _{0,3}	•••	0 \
	R _{1,0}	D _{1,1}	P _{1,2}	O _{1,3}	•••	O _{1,n}
Q =	O _{2,0}	R _{2,1}	D _{2,2}	P _{2,3}	•••	O _{2,n}
	:	÷	÷	÷	٠.	0 <i>O</i> _{1,n} <i>O</i> _{2,n} :
	\ 0	O _{n,1}	O _{n,2}	O _{n,3}	•••	$D_{n,n}$

where $P_{i,j}$ is a rectangular matrix of size $\binom{n}{i} \times \binom{n}{j}$, $R_{j,i}$ is a rectangular matrix of size $\binom{n}{j} \times \binom{n}{i}$, the symbol O_{ij} denotes a null matrix and D_{ii} are square diagonal matrices for i = 1, ..., n. Moreover,

$$-D_{i,i} = (e^{i}.R'_{i-1,i} + e^{i+1}.P'_{i,i+1}).I_{i,i}$$

where $I_{i,i}$ is the unit matrix of size *i*, e^{j} is the row vector of $\binom{n}{i}$ ones.



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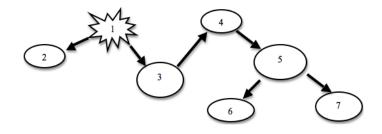
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Example

From all 2⁷ possible states only the following 14 are admissible:

(1,0,0,0,0,0,0), (1,1,0,0,0,0,0), (1,0,1,0,0,0,0), (1,1,1,0,0,0,0), (1,0,1,1,0,0,0), (1,1,1,1,0,0,0), (1,0,1,1,1,0,0), (1,0,1,1,1,0,0), (1,0,1,1,1,0,0), (1,1,1,1,1,0,0), (1,0,1,1,1,1,0,1), (1,0,1,1,1,1,1,0), (1,1,1,1,1,0,1), (1,0,1,1,1,1,1,1), (1,0,1,1,1,1,1).

The initial probability vector is equal to $\pi = (1, 0, 0, 0, 0, 0, 0)$.



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The transient intensity matrix \boldsymbol{S} has the form

(u	а	b	0	0	0	0	0	0	0	0	0	0	0/
0	-0	0 0	С	0	0	0	0	0	0	0	0	0	0
0	0	V	d	е	0	0	0	0	0	0	0	0	0
0	0	0	-1	f 0	f	0	0	0	0	0	0	0	0
0	0	0	0	W	g	h	0	0	0	0	0	0	0
0	0	0	0	0	_	i 0	i	0	0	0	0	0	0
0	0	0	0	0	0	х	j	k	0	0	0	0	0
0	0	0	0	0	0	0	-1	0	0	Ι	0	0	0
0	0	0	0	0	0	0	0	У	0	т	0	n	0
0	0	0	0	0	0	0	0	0	Ζ	0	р	q	0
0	0	0	0	0	0	0	0	0	0	-r	0	0	r
0	0	0	0	0	0	0	0	0	0	0	-s	0	s
0	0	0	0	0	0	0	0	0	0	0	0	—	tt
0/	0	0	0	0	0	0	0	0	0	0	0	0	0/

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There remains 19 unknown parameters which correspond to conditional transient intensities:

$$\begin{array}{lll} a=i(1\rightarrow 2|1); & g=i(1\rightarrow 2|1,3,4); & m=i(1\rightarrow 2|1,3,4,5,6); \\ b=i(1\rightarrow 3|1); & h=i(4\rightarrow 5|1,3,4); & n=i(5\rightarrow 7|1,3,4,5,6); \\ c=i(1\rightarrow 3|1,2); & i=i(4\rightarrow 5|1,2,3,4); & p=i(1\rightarrow 2|1,3,4,5,7); \\ d=i(1\rightarrow 2|1,3); & j=i(1\rightarrow 2|1,3,4,5); & q=i(5\rightarrow 6|1,3,4,5,7); \\ e=i(3\rightarrow 4|1,2); & k=i(5\rightarrow 6|1,3,4,5); & r=i(5\rightarrow 7|1,2,3,4,5,6); \\ f=i(3\rightarrow 4|1,3); & l=i(5\rightarrow 6|1,2,3,4,5); & s=i(5\rightarrow 6|1,2,3,4,5,7); \\ t=(1\rightarrow 2|1,3,4,5,6,7). \end{array}$$

In the case, where transitions between objects are independent of the previous path, we have

a = d = g = j = m = p = t, b = c, e = f, h = i, k = l = q = s, n = r. The whole system can be reduced to 6 unknown parameters a, b, e, h, k, n.



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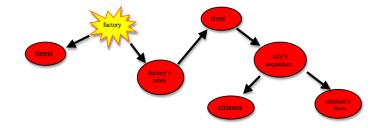
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One-way Model: The time of the total deterioration of the system is the time *T*, in which the process X(t) will attach the state ω^N .

During this time, the pursued events will pass through all objects of the system, the state ω^N is absorbing.



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Proposition

In the model 1, the time T to the system deterioration is a random variable, which has phase type probability distribution with initial probability vector $\pi = (\pi_1, \pi_2, ..., \pi_n, 0, 0, ..., 0)$ and upper triangular transition intensity matrix

$$\mathbf{S} = \begin{pmatrix} D_{1,1} & P_{1,2} & O_{1,3} & \cdots & O_{1,n-1} \\ O_{2,1} & D_{2,2} & P_{2,3} & \cdots & O_{2,n-1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ O_{n-1,1} & O_{n-1,2} & O_{n-1,3} & \cdots & D_{n-1,n-1} \end{pmatrix}$$



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For PH-distribution, we know all moments

$$E(T^k)=(-1)^kk!\pi S^{-k}e, \ k\in N.$$

where e is vector of ones.



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In model 2, the time T to the system deterioration is a random variable, which has phase type probability distribution with initial probability vector $\pi = (1, 0, ..., 0)$ and transition intensity matrix

$$S = \begin{pmatrix} D_{0,0} & P_{0,1} & O_{0,2} & \cdots & O_{0,n-1} \\ R_{1,0} & D_{1,1} & P_{1,2} & \cdots & O_{1,n-1} \\ 0_{2,0} & R_{2,1} & D_{2,2} & \cdots & O_{2,n-1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ O_{n-1,0} & O_{n-1,1} & O_{n-1,2} & \cdots & D_{n-1,n-1} \end{pmatrix}$$



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.

Particularly,

$$E(T) = -\pi S^{-1} e,$$

$$Var(T) = 2\pi S^{-2}e - (\pi S^{-1}e)^2.$$



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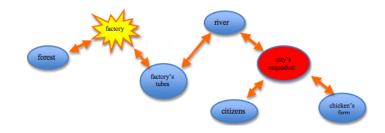
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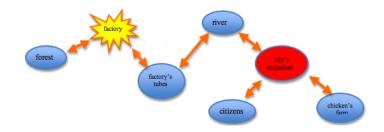
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For given *i* denote

$$\mathbf{S}_{\mathbf{0}}^{i} = \{\omega \in \Omega : \omega_{i} = \mathbf{0}\}$$

$$\mathbf{S}_{\mathbf{1}}^{i} = \{\omega \in \Omega : \omega_{i} = \mathbf{1}\}$$



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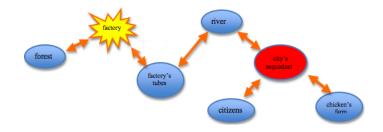
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For given *i* denote

 $S_0^i = \{ \omega \in \Omega : \omega_i = 0 \} - \text{the set of states of the system, in} \\ \text{which the } i\text{-th object is not affected} \\ S_1^i = \{ \omega \in \Omega : \omega_i = 1 \}$



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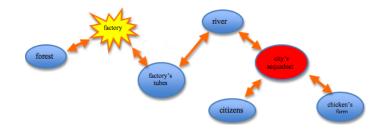
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For given *i* denote

$$\mathbf{S}_{\mathbf{0}}^{i} = \{\omega \in \Omega : \omega_{i} = \mathbf{0}\}$$

 $S_1^i = \{ \omega \in \Omega : \omega_i = 1 \}$ – the set of states, in which the *i*-th object is affected



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For given *i* denote

$$\mathbf{S}_{\mathbf{0}}^{i} = \{\omega \in \Omega : \omega_{i} = \mathbf{0}\}$$

$$\mathbf{S}_{\mathbf{1}}^{i} = \{\omega \in \Omega : \omega_{i} = \mathbf{1}\}$$

The sets S_0^i and S_1^i are disjoint and $S_0^i = \Omega - S_1^i$



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One-way Model

Reordering states in both of these sets in ascending order according to $|\omega|$, we can write the transition intensity matrix in the following block form:

 $\begin{array}{ccc} S_0^i & S_1^i \\ S_0^i & \begin{pmatrix} A & C \\ O & B \end{pmatrix} \end{array}$

where

A is upper triangular of the size $(2^{n-1} - 1) \times (2^{n-1} - 1)$, B is is upper triangular of the size $(2^{n-1} \times 2^{n-1})$,



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One-way Model

Reordering states in both of these sets in ascending order according to $|\omega|$, we can write the transition intensity matrix in the following block form:

 $\begin{array}{ccc} S_0' & S_1' \\ S_0' & \begin{pmatrix} A & C \\ O & B \end{pmatrix} \end{array}$

where

A is upper triangular of the size $(2^{n-1} - 1) \times (2^{n-1} - 1)$, B is is upper triangular of the size $(2^{n-1} \times 2^{n-1})$,

Repairable Model

After reordering of states we obtain the transition intensity matrix

$$\begin{array}{cccc}
S_0' & S_1' \\
S_0' & \begin{pmatrix} A & C \\
G & B \end{pmatrix}
\end{array}$$

where G is not null matrix.



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The first affect time T_i for an object *i* in the system is the random variable with the phase type probability distribution with representation (π^i , A).



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Proposition

The first affect time T_i for an object *i* in the system is the random variable with the phase type probability distribution with representation (π^i , A).

 $\pi^{i} = (\pi_{1}, \ldots, \pi_{i-1}, \pi_{i+1}, \ldots, \pi_{n-1}, 0, \ldots, 0, \pi_{i}),$



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$$\pi^{i} = (\pi_{1}, \ldots, \pi_{i-1}, \pi_{i+1}, \ldots, \pi_{n-1}, 0, \ldots, 0, \pi_{i}),$$

A is a matrix in the similar form as the matrix *S* in the previous Propositions: in **the One-way Model**

$$A = \begin{pmatrix} D_{1,1} & P_{1,2} & O_{1,3} & \cdots & O_{1,n-1} \\ O_{2,1} & D_{2,2} & P_{2,3} & \cdots & O_{2,n-1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ O_{n-1,1} & O_{n-1,2} & O_{n-1,3} & \cdots & D_{n-1,n-1} \end{pmatrix}$$

where

 $P_{i,j}$ is a rectangular matrix of size $\binom{n}{i} \times \binom{n}{j}$, O_{ij} is a null matrix, D_{ii} are square diagonal matrices for i = 1, ..., n. Moreover $D_{i,i} = -(e.P'_{i,i+1}.I_{i,i})$, where e is the row vector of $\binom{n}{j}$ ones and $I_{i,i}$ is the unit matrix of size i.



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$$\pi^{i} = (\pi_{1}, \ldots, \pi_{i-1}, \pi_{i+1}, \ldots, \pi_{n-1}, 0, \ldots, 0, \pi_{i}),$$

A is a matrix in the similar form as the matrix *S* in the previous Propositions: in **the Repairable Model**

$$A = \begin{pmatrix} D_{0,0} & P_{0,1} & O_{0,2} & \cdots & O_{0,n-1} \\ R_{1,0} & D_{1,1} & P_{1,2} & \cdots & O_{1,n-1} \\ 0_{2,0} & R_{2,1} & D_{2,2} & \cdots & O_{2,n-1} \\ \vdots & \vdots & \vdots & \ddots & \vdots \\ O_{n-1,0} & O_{n-1,1} & O_{n-1,2} & \cdots & D_{n-1,n-1} \end{pmatrix}$$

where $R_{j,i}$ are rectangular matrices of size $\binom{n}{j} \times \binom{n}{i}$, and $D_{i,i} = -(e^{i}.R'_{i-1,i} + e^{i+1}.P'_{i,i+1}).I_{i,i}$



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Under the assumptions of **model 2**, the process can be understood as renewal process. The renewal occurs when the process reaches the state ω^0 and the renewal period covers the time which the process needs to return to the state ω^0 again, if it started in ω^0 .



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For this purpose, we shall consider the state ω^0 in two manners:

- as a starting state ($\pi^0 = (1, 0, \dots, 0)$)
- as an absorption state after the process will move to it form another state.



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For this purpose, we shall consider the state ω^0 in two manners:

- as a starting state ($\pi^0 = (1, 0, \dots, 0)$)
- as an absorption state after the process will move to it form another state.
- ⇒ We can describe the distribution of renewal period using PH-distribution.



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Proposition

In the model 2, the distribution of renewal period can be described by PH-distribution with the representation (π^0 , W), where

	$(D_{0,0})$	P _{0,1}	<i>O</i> _{0,2}	O _{0,3}	•••	0 \
	O _{1,0}	$D_{1,1}$	P _{1,2}	O _{1,3}	•••	O _{1,n}
W =	O _{2,0}	R _{2,1}	D _{2,2}	P _{2,3}	•••	O _{2,n}
	÷	÷	÷	÷	٠.	0 <i>O</i> _{1,n} <i>O</i> _{2,n} :
	\ 0	O _{n,1}	O _{n,2}	O _{n,3}	•••	$D_{n,n}$

where $P_{i,j}$ is a rectangular matrix of size $\binom{n}{i} \times \binom{n}{j}$, the symbol O_{ij} denotes a null matrix and D_{ii} are square diagonal matrices for i = 1, ..., n. Moreover,

$$-D_{i,i} = e.P_{i,i+1}'.I_{i,i}$$

where $I_{i,i}$ is the unit matrix of size *i*, *e* is the row vector of $\binom{n}{i}$ ones.



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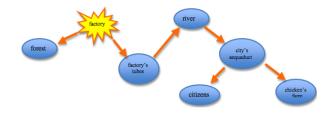
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There has been proposed the model for deterioration spreading in two variants:





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There has been proposed the model for deterioration spreading in two variants:

One-way Model (for nonrepairable systems)





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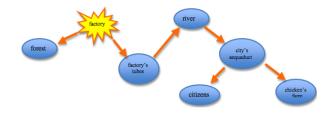
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Conclusions

There has been proposed the model for deterioration spreading in two variants:

- One-way Model (for nonrepairable systems)
- Repairable Model (when we permit repairs of system components)





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- There has been proposed the model for deterioration spreading in two variants:
 - One-way Model (for nonrepairable systems)
 - Repairable Model (when we permit repairs of system components)
- The model is supposed to be Markovian.





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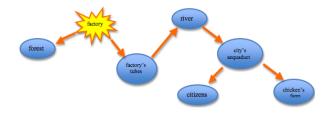
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Conclusions

- There has been proposed the model for deterioration spreading in two variants:
 - One-way Model (for nonrepairable systems)
 - Repairable Model (when we permit repairs of system components)
- The model is supposed to be Markovian.
- The first-passage time is modelled using PH distribution.





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Thanks for your attention.

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